# Reconstruction of a persistent random walk from exit time distributions 

Pak-Wing Fok ${ }^{1}$, Qunhui Han ${ }^{1}$, and Tom Chou ${ }^{2}$<br>${ }^{1}$ Dept. of Mathematical Sciences, University of Delaware, DE 19716<br>${ }^{2}$ Dept. of Biomathematics \& Dept. of Mathematics, UCLA, Los Angeles, CA 90095

March 4, 2013


#### Abstract

In this paper, we study the inverse problem of reconstructing the spatially dependent transition rate $F(x)$ of a one-dimensional Broadwell process from exit time distributions. In such a process, an advecting particle is assumed to undergo transitions between states with constant positive $(+v)$ and negative $(-v)$ velocities. The goal is to reconstruct the transition rate function $F(x)$ from the exit time distributions out of a finite interval. Using the associated backward equation, we compute the distribution of exit times and its Laplace transform, given a fixed starting position and velocity. We propose two methods (called ' $t$ ' and ' $s$ ') for finding $F(x)$. In both methods, we represent $F(x)$ as a linear combination of polynomials and repeatedly solve the backward equation to minimize the difference between its solution and given first exit time data. In the $t$-method we work in the time domain, using exit times directly and leveraging a novel series solution for the exit time distribution. In the $s$-method, we work with the Laplace-transformed equation and Laplace-transformed exit times. Noisy data is generated using a custom-designed algorithm to simulate the trajectories of a Broadwell process. In most cases we can find 4 coefficients to within $O\left(10^{-1}\right)$ accuracy from $O\left(10^{4}\right)$ exit times, with the $t$-method method slightly out-performing the $s$-method. We also explore the effectiveness of our algorithms for a fixed number of exit times under different advection speeds and find that optimal reconstruction occurs when $v=O(1)$.


Keywords: Random Walk, Inverse Problem, Broadwell, Telegrapher's equation

## 1 Introduction

Inverse problems arise in many applications such as medical imaging [1], high energy particle physics [21], and seismology [9]. Most of these applications involve measurement of waves at
the boundary of a domain; from this boundary data one may wish to reconstruct spatiallydependent properties within the domain such as the density and/or wave speed. However, there have been fewer successful applications where the underlying physics involves an intrinsic random process. The corresponding "boundary data" for stochastic inverse problems are probability fluxes or exit time distributions. The types of stochastic inverse problems we are concerned with in this paper involve inferring the parameters of a stochastic process from such a distribution.

One example of such a stochastic inverse problem arises in the reconstruction of bond potentials from rupture time distributions [10, 12, 17]. In force spectroscopy experiments, an increasing force is applied across a macromolecular bond until it ruptures. Because of thermal fluctuations, the rupture force is a random variable; thus the goal is to infer properties of the bond potential from the distribution of rupture forces. Stochastic inverse problems also commonly arise in diffuse optical tomography [1, 2]. In all these applications, the exit time distribution of a Brownian motion leaving a finite interval is measured, and one wishes to reconstruct the drift and/or diffusion function.

While Brownian motion is a canonical stochastic model, the inverse problem associated with Brownian motion is ill-posed $[4,11]$ and motivates the study of stochastic exit time problems based on other types of random walk. Ill-posedness is a trademark of many inverse problems. A problem is well-posed if a solution exists, is unique and depends continuously on the data. Otherwise the problem is ill-posed. At present, issues of existence and uniqueness of spatially dependent parameters for random walks are generally not well established, although some important results for Brownian motions can be found in [4].

In this paper, we generalize the study of Brownian inverse problems by focusing on a class of persistent random walk models called Broadwell processes [6, 7, 8, 16, 20]. In a Broadwell process, a particle randomly interconverts ("flips") between multiple states with each state associated with a particular velocity. The Broadwell process has the desirable property that it interpolates between a ballistic and diffusive motion [5, 22]: the time between transitions decreases as the flip rate increases, but increases as the flip rate decreases. The Broadwell model opens the analysis of the inverse problem for these two types of limiting processes; studying the inverse Broadwell problem may therefore provide insight into the important, fully diffusive problem. We specialize to the constant speed Broadwell process, assuming that transition probabilities that are spatially dependent and that the particle takes only two states associated with a positive and negative velocity. From the distribution of exit times out of a finite interval, our goal is to find the transition probability ("flip-rate") function.

Accurately simulating the exit times of a Brownian motion can be quite involved although reliable methods do exist: see $[14,15,19,25]$ for example. Nevertheless, one important advantage of studying the Broadwell model is the ease with which it can be simulated. Accurate simulations are critical for comparisons between reconstructed flip-rate functions and the underlying target functions that produced the exit data. When generating the exit time distribution from our simulations, the only source of error (besides round-off) stems from using a finite number of realizations.

The outline of this paper is as follows. In section A, we derive from first principles the Backward Kolmogorov Equation (BKE) for a general $K$-state Broadwell process. We then restrict our attention to the simple two-state case parameterized by a spatially dependent
(but state-symmetric) transition rate and a constant speed. We also state the inverse problem of reconstructing the flip-rate function from the exit time distributions. The associated optimization problem involves minimizing the distance between the solution of the BKE and the exit time data (derived from simulations or from the solution of the BKE with a given flip rate function). In section 2 we discuss the numerical aspects of our work. In particular, we present two reconstruction methods. The first involves minimizing the difference between the solution to the BKE and the target data in the time domain. The second involves minimizing the difference between the Laplace-transformed solution of the BKE and transformed exit time data. We also explain the simulation of Broadwell random walkers to test our reconstruction protocols. In Section 3 we present the results of our reconstruction using noisy data and compare the two methods. We show that for a finite number of exit times, the most reliable reconstruction of the flip rate occurs at an intermediate advection speed, no matter which method is used. In Section 4, we discuss general implications of our results and summarize our findings.

### 1.1 Two-state model and statement of inverse problem

A two-state Broadwell model describes a particle that can take one of two states. Initially, the particle is at position $x$ and in state $i \in\{1,2\}$. The particle advects within an interval $(-L / 2, L / 2)$ with velocity $+v$ if $i=1$ and $-v$ if $i=2$. While advecting, the particle may change state with probability $F(y) \mathrm{d} t$ within time interval $(t, t+\mathrm{d} t)$ where $y$ is the current particle position.

Let $w(t \mid x, 1,0) \equiv w_{1}(x, t)$ and $w(t \mid x, 2,0)=w_{2}(x, t)$ be the exit time distributions conditioned on the particle initially being at position $x$ and having a positive $(+v)$ and negative $(-v)$ velocity respectively. Then the exit time distributions satisfy

$$
\begin{align*}
\frac{\partial w_{1}}{\partial t} & =v \frac{\partial w_{1}}{\partial x}+F(x)\left(w_{2}-w_{1}\right)  \tag{1}\\
\frac{\partial w_{2}}{\partial t} & =-v \frac{\partial w_{2}}{\partial x}+F(x)\left(w_{1}-w_{2}\right) \tag{2}
\end{align*}
$$

subject to initial conditions

$$
\begin{equation*}
w_{1}(x, 0)=0, \quad w_{2}(x, 0)=0 \tag{3}
\end{equation*}
$$

and boundary conditions

$$
\begin{equation*}
w_{1}(x=L / 2, t)=\delta(t), \quad w_{2}(x=-L / 2, t)=\delta(t) \tag{4}
\end{equation*}
$$

A full derivation of the backward equation for the exit time distribution for a general $K$-state Broadwell process (of which (1)-(4) are a special case) can be found in Appendix A.

Given $F(x)$, one can solve eqs. (1)-(3) (the "forward problem") to find $w_{1}(x, t)$ and $w_{2}(x, t)$ for all $-L / 2<x<L / 2$ and $t>0$. Note that eqs. (1) and (2) constitute Kolmogorov's backward equations for the exit time distributions, and that solving this backward equation defines the forward problem for computing $w_{1}(x, t), w_{2}(x, t)$ from a known $F(x)$. However in
this paper we are interested in the inverse problem:
Problem Statement: Consider eqs. (1)-(3). Given a known, fixed $-L / 2<x_{0}<L / 2$, a known velocity $v>0$ and exit time distributions $w_{1}\left(x_{0}, t\right)$ and $w_{2}\left(x_{0}, t\right)$ for $t>0$, find $F(x) \in C(-L / 2, L / 2)$.

In practice, the exit time distributions could come from directly simulating a Broadwell process or from a single solution of the forward problem. For particles that initially advect with velocities $+v$ and $-v$, we refer to associated exit time distributions $w_{1, \text { data }}\left(x_{0}, t\right)$ and $w_{2, \text { data }}\left(x_{0}, t\right)$ respectively; note that $w_{1, \text { data }}\left(x_{0}, t\right)$ and $w_{2, \text { data }}\left(x_{0}, t\right)$ may or may not be noisy. We give details on how $w_{1, \text { data }}$ and $w_{1, \text { data }}$ are computed in Section 2.2.

Unfortunately, eqs. (1)-(3) are not useful in practice for inferring $F(x)$ because the solutions are highly singular. Related quantities that are more regular and whose governing equations are more amenable to numerical methods are the cumulative density functions (cdfs) and Laplace-transformed probability density functions. We now give explicit forms for these equations since we make frequent use of them later on.

The cdfs are related to the pdfs by $W_{1,2}(x, t)=\int_{0}^{t} w_{1,2}\left(x, t^{\prime}\right) \mathrm{d} t^{\prime}$. Therefore, upon integrating (1) and (2) in time, we find

$$
\begin{align*}
& \frac{\partial W_{1}}{\partial t}-v \frac{\partial W_{1}}{\partial x}=F(x)\left(W_{2}-W_{1}\right)  \tag{5}\\
& \frac{\partial W_{2}}{\partial t}+v \frac{\partial W_{2}}{\partial x}=F(x)\left(W_{1}-W_{2}\right) \tag{6}
\end{align*}
$$

subject to the boundary conditions

$$
\begin{equation*}
W_{1}(L / 2, t)=H(t), \quad W_{2}(-L / 2, t)=H(t) \tag{7}
\end{equation*}
$$

and initial conditions

$$
\begin{equation*}
W_{1}(x, 0)=0, \quad W_{2}(x, 0)=0 . \tag{8}
\end{equation*}
$$

In eqs. (7), $H(t)$ is the Heaviside step function satisfying $H(t)=1$ if $t>0$ and $H(t)=0$ if $t \leq 0$. The corresponding inverse problem is to find $F(x) \in C(-L / 2, L / 2)$ given $-L / 2<$ $x_{0}<L / 2, v$ and $W_{j, \text { data }}\left(x_{0}, t\right)=\int_{0}^{t} w_{j, \text { data }}\left(x_{0}, t^{\prime}\right) \mathrm{d} t^{\prime}$ for $t>0$ and $j=1,2$. Alternatively, we can also take Laplace transforms of (1) and (2) to find

$$
\begin{align*}
& s \tilde{w}_{1}(x, s)=v \frac{\partial \tilde{w}_{1}}{\partial x}+F(x)\left(\tilde{w}_{2}-\tilde{w}_{1}\right)  \tag{9}\\
& s \tilde{w}_{2}(x, s)=-v \frac{\partial \tilde{w}_{2}}{\partial x}+F(x)\left(\tilde{w}_{1}-\tilde{w}_{2}\right) \tag{10}
\end{align*}
$$

subject to boundary conditions

$$
\begin{equation*}
\tilde{w}_{1}(x=L / 2, s)=1, \quad \tilde{w}_{2}(x=-L / 2, s)=1 . \tag{11}
\end{equation*}
$$

The corresponding inverse problem is to find $F(x) \in C(-L / 2, L / 2)$ given $-L / 2<x_{0}<L / 2$, $v$ and $\tilde{w}_{j, \text { data }}\left(x_{0}, s\right)$ for $s>0$ and $j=1,2$.

## 2 Algorithms for Reconstruction

This section is divided into 3 parts. First, we give details on solving the forward problems (9)-(11) and (5)-(8). Then we explain how to generate noisy data by directly simulating a Broadwell process with spatially dependent flip rate. Finally, we discuss a projection method to solve the inverse problem.

### 2.1 Solution to the forward problems

Solution to (9)-(11): Our method for finding $\tilde{w}_{1}\left(x_{0}, s\right)$ and $\tilde{w}_{2}\left(x_{0}, s\right)$ from eqs. (9)-(11) is based on solving the boundary value problem using a pseudospectral method [26] for different values of $s \geq 0$ : see Algorithm 1. The solutions $\tilde{w}_{1,2}\left(x_{0}, s\right)$ are always infinitely differentiable, monotonically decreasing functions in $s$ that $\rightarrow 0$ as $s \rightarrow \infty$.

```
Algorithm 1 Algorithm for solving the forward problem (9)-(11).
    Require: flip rate function \(F(x)\), velocity \(v>0\), interval size \(L\), starting position \(-L / 2<\)
    \(x_{0}<L / 2\) and integer \(N \gg 1\).
    for \(i=1,2, \ldots, N\) do
        let \(\xi_{i}=(i-1) / N\) and \(s_{i}=\xi_{i} /\left(1-\xi_{i}\right)\).
        With \(s=s_{i}\), solve (9)-(11) using a pseudospectral discretization [26] in \(x\).
        Interpolate the solution at \(x=x_{0}\) to find \(\tilde{w}_{1}\left(x_{0}, s_{i}\right)\) and \(\tilde{w}_{2}\left(x_{0}, s_{i}\right)\).
    end for
    Output: Laplace Transformed exit time distributions \(\tilde{w}_{1}\left(x_{0}, s_{i}\right)\) and \(\tilde{w}_{2}\left(x_{0}, s_{i}\right), i=\)
    \(1, \ldots, N\).
```

Solution to (5)-(8): In contrast to $\tilde{w}_{1,2}\left(x_{0}, s\right)$, the solutions $W_{1,2}(x, t)$ contain jump discontinuities that propagate into the domain of solution with velocity $\pm v$ : the jump discontinuity in $W_{1}(x, t)\left(W_{2}(x, t)\right)$ propagates along the characteristic line $t=-x / v+L /(2 v)$ $(t=+x / v+L /(2 v))$. This behavior in the singularities is illustrated by the following theorem which uses an eigenfunction expansion to construct an explicit solution to (5)-(8).

Theorem 1 (Series solution to the forward problem (5)-(8)). For $0<t<L / v$, the solution to (5)-(8) is

$$
\begin{align*}
W_{1}(x, t) & =a_{1}(x) H[t+x / v-L /(2 v)]+Z_{1}(x, t),  \tag{12}\\
W_{2}(x, t) & =a_{2}(x) H[t-x / v-L /(2 v)]+Z_{2}(x, t), \tag{13}
\end{align*}
$$

where $H[\cdot]$ is the heaviside step function and $Z_{1,2}(x, t)$ are continuous functions given by the series

$$
\begin{equation*}
\mathbf{Z}(x, t)=-\sum_{m=1}^{\infty} \frac{\mathbf{u}_{m}(x)}{s_{m} D_{m}}\left[h_{m}^{(1)}(t)+h_{m}^{(2)}(t)\right] \tag{14}
\end{equation*}
$$

where:

$$
\begin{align*}
h_{m}^{(1)}(t) & =\int_{-L / 2}^{-L / 2+v t} p_{m}^{(1) *}(y) a_{2}(y) F(y)\left(1-e^{\left(t-\frac{y}{v}-\frac{L}{2 v}\right) s_{m}}\right) d y,  \tag{15}\\
h_{m}^{(2)}(t) & =\int_{L / 2-v t}^{L / 2} p_{m}^{(2) *}(y) a_{1}(y) F(y)\left(1-e^{\left(t+\frac{y}{v}-\frac{L}{2 v}\right) s_{m}}\right) d y,  \tag{16}\\
a_{1}(x) & =\exp \left[-\frac{1}{v} \int_{x}^{L / 2} F\left(x^{\prime}\right) d x^{\prime}\right],  \tag{17}\\
a_{2}(x) & =\exp \left[-\frac{1}{v} \int_{-L / 2}^{x} F\left(x^{\prime}\right) d x^{\prime}\right],  \tag{18}\\
D_{m} & =\left\langle\mathbf{p}_{m}(x), \mathbf{u}_{m}(x)\right\rangle=\int_{-L / 2}^{L / 2} \mathbf{p}_{m}^{*}(x) \mathbf{u}_{m}(x) d x . \tag{19}
\end{align*}
$$

In (14), $s_{m} \in \mathbb{C}$ and $\mathbf{u}_{m}(x) \in \mathbb{C}^{2}$ are the eigenvalues and eigenfunctions of $\mathbf{A}$ where

$$
\mathbf{A}\binom{u_{1}}{u_{2}}=\left\{\left[\begin{array}{cc}
v \frac{\mathrm{~d}}{\mathrm{~d} x} & 0  \tag{20}\\
0 & -v \frac{\mathrm{~d}}{\mathrm{~d} x}
\end{array}\right]+F(x)\left[\begin{array}{rr}
-1 & 1 \\
1 & -1
\end{array}\right]\right\}\binom{u_{1}(x)}{u_{2}(x)}
$$

along with the boundary conditions $u_{1}(L / 2)=u_{2}(-L / 2)=0 . \mathbf{p}_{m}(x)=\left[p_{m}^{(1)}(x), p_{m}^{(2)}(x)\right]^{T}$ are the eigenfunctions of the adjoint operator $\mathbf{A}^{*}$.

We now discuss the behavior of the solutions $W_{1,2}(x, t)$ in light of eqs. (12)-(13) and defer the proof of the theorem to the end of this section. From (12) and (13) it is clear that discontinuities in the boundary conditions (7) propogate into the interior. In Figure 1(a), $W_{1}$ is discontinuous on the diagonal line separating $A, C$ and $B, D$ while $W_{2}$ is discontinuous on the line separating $A, D$ and $B, C$. Because the hyperbolic system (5)-(8) has a finite wave speed $v>0$, region $C$ is outside the region of influence of the disturbances originating at $(x, t)=$ $(L / 2,0)$ and $(-L / 2,0)$ and we expect that $W_{1}(x, t)=W_{2}(x, t)=0$ in $C$. This behavior is confirmed in Figure 1(b) which shows the cumulative distribution functions $W_{1,2}(x=$ $-L / 4, t)$ calculated using (12)-(14). The function $W_{1}(-L / 4, t)$ has a discontinuous derivative at $t_{1}=L /(4 v)$ and a jump discontinuity at $t_{2}=3 L /(4 v)$ while $W_{2}(-L / 4, t)$ has a jump discontinuity at $t_{1}$ and a discontinuous derivative at $t_{2}$. Figure 1(c) shows cdfs evaluated at $x_{0}=L / 3$. The inset shows associated Laplace-transformed probability density functions $\tilde{w}_{1}(L / 3, s)$ and $\tilde{w}_{2}(L / 3, s)$, found by solving eqs. (9)-(11)) using Algorithm 1. Cumulative density functions from Monte-Carlo simulations are superposed to validate our numerical method; details of how these simulations are performed are described in section 2.2.

The expansions (12)-(13) in Theorem 1 are commonly used to analyze seismic waves $[9,23,24]$ and form the basis of our numerical method for the forward problem in $t$ : see Algorithm 2. Numerically, the cdfs $W_{1,2}$ are computed by taking a finite number of terms in (14) and adding on a step discontinuity at $t=\mp x / v+L /(2 v)$ with strength given by (17) and (18). A pseudo-spectral collocation method on a Chebyshev grid was used to find the eigenvectors $\mathbf{u}_{j}$ and Clenshaw-Curtis quadrature [26] was used to quickly evaluate the

$$
\begin{aligned}
& \text { Algorithm } 2 \text { Algorithm for solving the forward problem (5)-(8). The same symbol is used } \\
& \text { to refer to quantities in (14)-(19) as well as their numerical approximations. For example, } \mathbf{A} \\
& \text { refers to the differential operator as well as its matrix approximation. } \\
& \hline \text { 1: Require: A target flip rate } F(x) \text {, velocity } v>0 \text {, an integer } N \text {, a starting position } \\
& -L / 2<x_{0}<L / 2 \text { and a discretization of the interval }[-L / 2, L / 2] \text {, }\left\{\chi_{0}, \chi_{1}, \ldots, \chi_{n}\right\} . \\
& \text { 2: Discretize the differential operators } \mathbf{A} \text { and the adjoint } \mathbf{A}^{*} \text { where } \\
& \qquad \mathbf{A}\binom{u_{1}}{u_{2}}=\left\{\left[\begin{array}{cc}
v \frac{\mathrm{~d}}{\mathrm{~d} x} & 0 \\
0 & -v \frac{\mathrm{~d}}{\mathrm{~d} x}
\end{array}\right]+F(x)\left[\begin{array}{rr}
-1 & 1 \\
1 & -1
\end{array}\right]\right\}\binom{u_{1}(x)}{u_{2}(x)}, \\
& \qquad \mathbf{A}^{*}\binom{p_{1}}{p_{2}}=\left\{\left[\begin{array}{rr}
-v \frac{\mathrm{~d}}{\mathrm{~d} x} & 0 \\
0 & v \frac{\mathrm{~d}}{\mathrm{~d} x}
\end{array}\right]+F(x)\left[\begin{array}{rr}
-1 & 1 \\
1 & -1
\end{array}\right]\right\}\binom{p_{1}(x)}{p_{2}(x)} .
\end{aligned}
$$

Note that A must account for the boundary conditions $u_{1}(L / 2)=0$ and $u_{2}(-L / 2)=0$ respectively and $\mathbf{A}^{*}$ must account for the adjoint boundary conditions $p_{1}(-L / 2)=0$ and $p_{2}(L / 2)=0$.
3: Compute $s_{1}, \ldots, s_{N}$, the first $N$ complex eigenvalues of $\mathbf{A}$ with smallest absolute value.
4: Compute the corresponding $N$ eigenvectors of $\mathbf{u}_{1}, \ldots, \mathbf{u}_{N}$ of $\mathbf{A}$ and $\mathbf{p}_{1}, \ldots, \mathbf{p}_{N}$ of $\mathbf{A}^{*}$.
5: Compute the inner products $D_{m}=\int_{-L / 2}^{L / 2} \mathbf{p}_{m}^{*}(x) \mathbf{u}_{m}(x) \mathrm{d} x$ for $m=1, \ldots, N$.
6: Compute the functions $h_{m}^{(1)}(t)$ and $h_{m}^{(2)}(t)$ in (15) and (16) for $m=1, \ldots, N$.
7: Compute $a_{1}(x)$ and $a_{2}(x)$ in (17) and (18).
8: Compute

$$
\begin{equation*}
\mathbf{Z}\left(x_{0}, t\right)=-\sum_{m=1}^{N} \frac{\mathbf{u}_{m}\left(x_{0}\right)}{s_{m} D_{m}}\left[h_{m}^{(1)}(t)+h_{m}^{(2)}(t)\right], \tag{21}
\end{equation*}
$$

as the $N$-term approximation to (14). If $x_{0}$ does not coincide with a grid point $\chi_{j}$, use interpolation to find $\mathbf{u}_{m}\left(x_{0}\right)$.
9: Compute $W_{1}\left(x_{0}, t\right)$ and $W_{2}\left(x_{0}, t\right)$ by adding discontinuities of strength $a_{1}\left(x_{0}\right)$ and $a_{2}\left(x_{0}\right)$ at $t=-x_{0} / v+L /(2 v)$ and $t=x_{0} / v+L /(2 v)$ respectively to $Z_{1}$ and $Z_{2}$ : see eqs. (12) and (13).
10: Output $W_{1}\left(x_{0}, t\right)$ and $W_{2}\left(x_{0}, t\right)$.


Figure 1: (a) Propagation of discontinuities of (1)-(3) in the $x$ - $t$ plane. The solution $w_{1}(x, t)$ is formally infinite on $x=L / 2-v t$ while $w_{2}(x, t)$ is infinite on $x=-L / 2+v t$. These singularities give rise to discontinuities in $W_{1}$ and $W_{2}$ that can be seen in (b,c). (b) Numerical computation of cumulative probability densities (cdfs) $W_{1,2}\left(x_{0}=-L / 4, t\right)$ and auxiliary functions $Z_{1,2}\left(x_{0}=-L / 4, t\right)$ computed through (14) and Algorithm 2 using using 101 Chebyshev grid points and 51 eigenfunctions $\mathbf{u}_{1}, \mathbf{u}_{2}, \ldots, \mathbf{u}_{51}$. (c) Numerical computation of cdfs $W_{1,2}\left(x_{0}=L / 3, t\right)$ (solid) along with results from Monte Carlo simulations (diamond). Inset shows Laplace-transformed probability densities $\tilde{w}_{1,2}\left(x_{0}=L / 3, s\right)$. Common parameters in (b,c) are $v=1, F(y)=1+y$ and $L=1$.
integrals (15) and (16) for $0<t<L / v$. The strength of this numerical method is that no integration in time is required to find $W_{1,2}\left(x_{0}, t\right)$ and the method allows quick evaluation of the cdfs at one fixed value of $x=x_{0}$. Its weakness is that many terms are usually required $(\gtrsim 100)$ in the expansion to obtain accurate results when $x_{0}$ is close to $\pm L / 2$. Furthermore we found that when $x_{0}= \pm L / 2$, the expansion (14) converged to a discontinuous function, giving $W_{1,2}\left(x_{0}, t\right)>1$ as $t \rightarrow \infty$; hence the properties of the series (14) still require further investigation at the domain boundaries.

Another important reason for separating out $W_{1}$ and $W_{2}$ into continuous and discontinuous components is to avoid Gibbs oscillations when solving for $\mathbf{Z}(x, t)$ as superpositions of eigenfunctions $\mathbf{u}_{n}$. These oscillations would introduce large errors into the solution to the forward problem (5)-(8) and therefore hinder the solution of the inverse problem.

Proof of Theorem 1. Upon substituting (12), (13) into (5), (6), we find that $Z_{1,2}(x, t)$ satisfy

$$
\begin{align*}
\frac{\partial Z_{1}}{\partial t}-v \frac{\partial Z_{1}}{\partial x}-F(x)\left(Z_{2}-Z_{1}\right) & =a_{2}(x) F(x) H[t-x / v-L /(2 v)]  \tag{22}\\
\frac{\partial Z_{2}}{\partial t}+v \frac{\partial Z_{2}}{\partial x}-F(x)\left(Z_{1}-Z_{2}\right) & =a_{1}(x) F(x) H[t+x / v-L /(2 v)] \tag{23}
\end{align*}
$$

subject to the homogeneous boundary conditions $Z_{1}(L / 2, t)=0, Z_{2}(-L / 2, t)=0$ and initial conditions $Z_{1}(x, 0)=0, Z_{2}(x, 0)=0$ and $a_{1,2}$ are defined by (17) and (18). We now find a
series representation for $Z_{1,2}(x, t)$. After taking Laplace transforms of (22) and (23), we find that $\tilde{\mathbf{Z}}(x, s)=\left[\tilde{Z}_{1}(x, s), \tilde{Z}_{2}(x, s)\right]^{T}$ satisfies

$$
\begin{aligned}
(\mathbf{A}-s \mathbf{I}) \tilde{\mathbf{Z}}(x, s) & =-\frac{\tilde{\mathbf{N}}(x, s)}{s} \\
\tilde{\mathbf{N}}(x, s) & =F(x)\left[\begin{array}{l}
a_{2}(x) e^{-\left(\frac{L}{2 v}+\frac{x}{v}\right) s} \\
a_{1}(x) e^{-\left(\frac{L}{2 v}-\frac{x}{v}\right) s}
\end{array}\right] .
\end{aligned}
$$

with boundary conditions $\tilde{Z}_{1}(L / 2, s)=0, \tilde{Z}_{2}(-L / 2, s)=0$. Eq. (24) has a solution of the form

$$
\begin{equation*}
\tilde{\mathbf{Z}}(x, s)=\sum_{n=1}^{\infty} \frac{c_{n}(s) \mathbf{u}_{n}(x)}{s-s_{n}} \tag{24}
\end{equation*}
$$

where the vector eigenfunctions $\mathbf{u}_{n}(x) \in \mathbb{C}^{2}$ satisfy $\mathbf{A} \mathbf{u}_{n}=s_{n} \mathbf{u}_{n}$ for eigenvalues $s_{n} \in \mathbb{C}$. As an aside, when $F(x)=F_{0}$ is a constant, one can show that the eigenfunctions are proportional to $\left[-\left(\alpha^{2}+\lambda_{n}^{2}\right)^{1 / 2} \sinh \lambda_{n}(x+1 / 2)+\lambda_{n} \cosh \lambda_{n}(x+1 / 2), \alpha \sinh \lambda_{n}(x+1 / 2)\right]^{T}$ with $\alpha \equiv F_{0} L / v$, the $\lambda_{n} \in \mathbb{C}$ satisfy the transcendental equation $-\left(\alpha^{2}+\lambda_{n}^{2}\right)^{1 / 2} \tanh \lambda_{n}+\lambda_{n}=0$, and the eigenvalues are given by $s_{n}=-\alpha-\left(\alpha^{2}+\lambda_{n}^{2}\right)^{1 / 2}$.

Recall that if $\left\{\mathbf{u}_{n}\right\}$ are the eigenfunctions of $\mathbf{A}$ and $\left\{\mathbf{p}_{m}\right\}$ are the eigenfunctions of the adjoint operator $\mathbf{A}^{*}$, then $\left\langle\mathbf{p}_{m}, \mathbf{u}_{n}\right\rangle=0$ unless $m=n$. Substituting (24) into (24), leftmultiplying both sides of by $\mathbf{p}_{m}^{*}$ and integrating, we find that

$$
c_{n}(s)=\frac{\left\langle\mathbf{p}_{m}(x), \tilde{\mathbf{N}}(x, s)\right\rangle}{s D_{m}}
$$

where $D_{m}$ is defined by (19). (One cannot obtain $c_{n}$ by invoking orthogonality of $\left\{\mathbf{u}_{n}\right\}$ since A is not self-adjoint.) We now take the inverse Laplace transform of (24) and switch the order of integration to obtain the continuous parts of the cumulative density functions (cdfs):

$$
\begin{aligned}
\mathbf{Z}(x, t)= & \sum_{m=1}^{\infty} \frac{\mathbf{u}_{m}(x)}{D_{m}}\left\{\int_{-L / 2}^{L / 2} \mathrm{~d} y p_{m}^{(1) *}(y) a_{2}(y) F(y) \int_{\gamma-i \infty}^{\gamma+i \infty} \frac{\mathrm{~d} s}{2 \pi i} \frac{e^{\left(t-\frac{y}{v}-\frac{L}{2 v}\right) s}}{s\left(s-s_{m}\right)}+\right. \\
& \left.\int_{-L / 2}^{L / 2} \mathrm{~d} y p_{m}^{(2) *}(y) a_{1}(y) F(y) \int_{\gamma-i \infty}^{\gamma+i \infty} \frac{\mathrm{~d} s}{2 \pi i} \frac{e^{\left(t+\frac{y}{v}-\frac{L}{2 v}\right) s}}{s\left(s-s_{m}\right)}\right\}, \quad \operatorname{Re} \gamma>0 \\
= & -\sum_{m=1}^{\infty} \frac{\mathbf{u}_{m}(x)}{s_{m} D_{m}}\left[h_{m}^{(1)}(t)+h_{m}^{(2)}(t)\right] .
\end{aligned}
$$

where $h_{m}^{(1)}(t)$ and $h_{m}^{(2)}(t)$ are given by eqs. (15) and (16) respectively.

### 2.2 Monte-Carlo simulation using Rejection-Acceptance

We now give details of our Monte-Carlo method in Algorithm 3. This method can be used to simulate a broadwell particle with spatially dependent velocity, even though for our inverse
problem, the particles always have constant velocity. The method is based on the RejectionAcceptance method [3], a common method for drawing random variables from a pdf whose functional form is known, but non-standard. We note four important points about the algorithm.

1. The algorithm samples from $w_{1}\left(y_{0}, t\right)$ or $w_{2}\left(y_{0}, t\right)$ depending on the initial velocity: see eqs. (27).
2. The algorithm generates random variables for the time periods in between the state transitions $\theta$ (the "flip times"). For a Broadwell process with a constant transition rate, the flip times are exponentially distributed. For a spatially-dependent transition rate $F(y)$, the flip time $\theta \equiv t_{j+1}-t_{j}$ is distributed according to

$$
\begin{align*}
\theta \sim Q(t) & \equiv F[y(t)] \exp [-p(t)], \\
p(t) & =\int_{0}^{t} F\left[y\left(t^{\prime}\right)\right] \mathrm{d} t^{\prime}, \tag{25}
\end{align*}
$$

where the position of the particle satisfies $\mathrm{d} y(t) / \mathrm{d} t=v(y)$. We sample from $Q(t)$ using a Rejection-Acceptance method [3]: suppose there exist constants $F_{\min }$ and $F_{\max }$ satisfying $0<F_{\min } \leq F(y) \leq F_{\max }<\infty$ for $-L / 2<y<L / 2$. Then

$$
Q(t) \leq C F_{\min } \exp \left(-F_{\min } t\right) \equiv P(t)
$$

where $C=F_{\max } / F_{\min }$ and so an exponential distribution can be used as an envelope function.
3. Once the flip time $\theta$ is generated, the flip position $y_{j+1}$ can be found by solving $\int_{y_{j}}^{y_{j+1}} v^{-1}(y) \mathrm{d} y=\theta$. This integral could be expensive to calculate if it has to be done many times. Also, every evaluation of $Q(t)$ requires computing the integral $p(t)$ in (25). Both of these issues are handled simultaneously in our algorithm through the solution of the pair of ordinary differential equations (26)-(??). For the special case where $v$ is a constant, (26) and (27) should be replaced with $\mathrm{d} p / \mathrm{d} t=F\left(y_{j}+v t\right)$ and $y(t)=y_{j}+v t$.
4. When solving the differential equations (26) and (27), $F$ may have to be evaluated outside of the interval $[-L / 2, L / 2]$. Because the form of $F(y)$ outside $[-L / 2, L / 2]$ does not affect the exit time, we simply take $F(y)=F(L / 2)$ for $y>L / 2$ and $F(y)=$ $F(-L / 2)$ for $y<-L / 2$.

Figure 2 shows the probability density functions $w_{1}(0, t)$ generated by the algorithm for two different $F(x)$ when $v(x)=$ constant. By definition, $w_{1}$ is the exit time density for a particle that initially has velocity $v>0$. Therefore the solution $w_{1}(x, t)$ in (1) contains delta functions that correspond to an immediate particle exit at a time $t_{c} \equiv(L / 2-x) / v$. The reason is if $t$ is the particle exit time and $\theta$ is the time before the first state transition, then

$$
\begin{equation*}
P\left(t=t_{c}\right)=P\left(\theta \geq t_{c}\right)=\int_{t_{c}}^{\infty} Q\left(t^{\prime}\right) \mathrm{d} t^{\prime}>0 . \tag{28}
\end{equation*}
$$

```
Algorithm 3 Generating exit times from a Broadwell process
    Require: an interval size \(L\), a starting position \(-L / 2 \leq y_{0} \leq L / 2\), functions \(F(y), v(y) \in\)
    \(C(-\infty,+\infty), F_{\min }, F_{\max }>0\) where \(F_{\min } \leq F(y) \leq F_{\max }\) for \(y \in[-L / 2, L / 2]\).
    Let \(P(t) \equiv F_{\max } \exp \left(-F_{\min } t\right)\).
    Set \(j=0\) and \(t_{j}=0\).
    while \(-L / 2<y_{j}<L / 2\) do
        Draw \(\theta \sim \exp \left(F_{\text {min }}\right)\)
        Compute \(p(\theta)\) and \(y(\theta)\) by numerically solving
\(\frac{\mathrm{d} p(t)}{\mathrm{d} t}=F(y)\),
\(\frac{\mathrm{d} y(t)}{\mathrm{d} t}=\left\{\begin{array}{rr}(-1)^{j} v(y), & \text { \{for positive velocity at } t=0\} \\ (-1)^{j+1} v(y), & \text { \{for negative velocity at } t=0\}\end{array}\right.\)
on \(t \in[0, \theta]\), subject to initial conditions \(p(0)=0\) and \(y(0)=y_{j}\).
7: \(\quad \operatorname{Set} Q(\theta)=F[y(\theta)] \exp [-p(\theta)]\)
Draw \(\rho \sim U(0,1)\)
if \(\rho<Q(\theta) / P(\theta)\) then
\(j \leftarrow j+1\) \{acceptance\}
Set \(y_{j}=y(\theta)\) and \(t_{j}=t_{j-1}+\theta\)
Goto 4
else
Goto 5 \{rejection\}
end if
end while\{Particle has left interval\}
if \(y_{j}>L / 2\) then
Output the exit time as \(t_{j}+\int_{y_{j}}^{L / 2} v^{-1}(y) \mathrm{d} y\).
else
Output the exit time as \(t_{j}+\int_{y_{j}}^{-L / 2}\left(-v^{-1}(y)\right) \mathrm{d} y\).
end if
```



Figure 2: Simulated exit times of a Broadwell process. Although probability densities are shown here, cumulative density functions are used to infer the flip rate function. (a) $F(x)=$ $10 x^{3}+5 e^{x}+1, v=1 / 2$ (b) $F(x)=1+x^{2}, v=1$. Insets show Laplace-transformed data. Solid line: solution to (9), (10). Red diamonds: Laplace transform of histogram data. The number of realizations was $N=40,000$ in each case.

Hence the probability distribution of the exit times will always contain point masses (delta functions) of probability located at $t=t_{c}$ [20]. In Fig. 2(b) a numerical approximation of this delta function can be seen. The height of this "spike" is controlled by the size of the bins used when creating the histogram and becomes unbounded as the bin size tends to zero and the number of trials tends to infinity. These delta distributions are always present in the exact solution but they may not always be visible in the numerical solution if the number of trials is small or the bin size is large; see Fig. 2(a) for example.

### 2.2.1 Generation of noisy distributions from a finite number of exit times

We always use Algorithm 3 to generate two sets of $N$ exit times $\left\{\tau_{j}^{(1)}\right\}$ and $\left\{\tau_{j}^{(2)}\right\}$. With this notation, $\tau_{j}^{(i)}(1 \leq j \leq N, i=1,2)$ is the $j$ th exit time conditioned on the particle having initial velocity $(-1)^{i+1} v$. Assuming that $\left\{\tau_{j}^{(1)}\right\}$ and $\left\{\tau_{j}^{(2)}\right\}$ are sorted in ascending order, noisy cumulative densities $W_{1, \text { data }}(t)$ and $W_{2, \text { data }}(t)$ are computed as

$$
W_{i, \text { data }}(t)=\left\{\begin{align*}
0 & \text { if } t<\tau_{1}^{(i)},  \tag{29}\\
\frac{m}{N} & \text { if } \tau_{1}^{(i)}<t<\tau_{N}^{(i)}, \\
1 & \text { if } t \geq \tau_{N}^{(i)},
\end{align*}\right.
$$

where $m$ is the unique index satisfying $\tau_{m}^{(i)}<t<\tau_{m+1}^{(i)}$.

The noisy Laplace transform of a finite number of exit times is calculated through

$$
\tilde{w}_{j, \text { data }}(s)=\left\{\begin{array}{cc}
\int_{0}^{\infty} e^{-s t} \frac{\mathrm{~d} W_{j, \text { data }}(t)}{\mathrm{d} t} \mathrm{~d} t=s \tilde{W}_{j, \text { data }}(s), & s>0,  \tag{30}\\
1, & s=0,
\end{array}\right.
$$

where

$$
\begin{equation*}
\tilde{W}_{j, \text { data }}(s)=\int_{0}^{\infty} e^{-s t} W_{j, \text { data }}(t) \mathrm{d} t=\int_{0}^{1} e^{-s t[\eta]} W_{j, \text { data }}(t[\eta]) \frac{\mathrm{d} \eta}{(1-\eta)^{2}}, \tag{31}
\end{equation*}
$$

where $j=1,2$ and $t[\eta]=\eta /(1-\eta)$. We avoid "binning" when calculating $\tilde{w}_{j, \text { data }}$, since this introduces a corresponding discretization error. The integral in (31) can be calculated using the trapezium rule on equally spaced abscissae in $\eta$.

### 2.3 Projection method to solve inverse problem

We now briefly describe our algorithms for reconstructing the flip rate function $F(x)$ from the two distributions of exit times $w_{1, \text { data }}(t)$ and $w_{2, \text { data }}(t)$. These distributions can come from simulating the Broadwell process directly through Algorithm 3 or through a one-time solution of the forward problems (9)-(11) or (5)-(8). We implement two related algorithms. The first method uses the exit time data directly ( $t$-method) and the second method uses Laplace-transformed exit time data ( $s$-method). Pseudocode for the two methods is given in Algorithms 4 and 5.

In both methods, we represent the trial flip rate function $F_{M}(x)$ and the target flip rate function $F^{*}(x)$ as linear combination of Legendre polynomials on $[-L / 2, L / 2]$ :

$$
\begin{equation*}
F_{M}(x)=\sum_{j=0}^{M-1} a_{j} \phi_{j}(x), \tag{32}
\end{equation*}
$$

For example, $\phi_{0}(x)=1, \phi_{1}(x)=2 x / L, \phi_{2}(x)=6\left(\frac{x}{L}\right)^{2}-\frac{1}{2}$.
Our aim is to find coefficients $a_{1}, \ldots, a_{M}$ to minimize the objective functions for the $t$-method and $s$-methods, $\Pi_{1}$ and $\Pi_{2}$ respectively. These take the form

$$
\begin{aligned}
& \Pi_{1}(\mathbf{a})=\int_{0}^{L / v}\left|W_{1}\left(x_{0}, t ; \mathbf{a}\right)-W_{1, \text { data }}\left(x_{0}, t\right)\right|^{2} \mathrm{~d} t+\int_{0}^{L / v}\left|W_{2}\left(x_{0}, t ; \mathbf{a}\right)-W_{2, \text { data }}\left(x_{0}, t\right)\right|^{2} \mathrm{~d} t \\
& \Pi_{2}(\mathbf{a})=\int_{0}^{\infty}\left|\tilde{w}_{1}\left(x_{0}, s ; \mathbf{a}\right)-\tilde{w}_{1, \text { data }}\left(x_{0}, s\right)\right|^{2} \mathrm{~d} s+\int_{0}^{\infty}\left|\tilde{w}_{2}\left(x_{0}, s ; \mathbf{a}\right)-\tilde{w}_{2, \text { data }}\left(x_{0}, s\right)\right|^{2} \mathrm{~d} s
\end{aligned}
$$

The data sets $W_{j, \text { data }}\left(x_{0}, t\right), \tilde{w}_{j, \text { data }}\left(x_{0}, s\right)$ associated with $F^{*}(x)$ can be computed from individual exit times using (29) and (30) respectively.

Minimization of $\Pi_{1}$ and $\Pi_{2}$ with respect to a was performed using the Matlab routines fminunc.m and lsqnonlin.m for the $t$ and $s$ methods respectively, with the tolerances TolFun and TolX set to $10^{-14}$. The initial guess for the coefficients was always $a_{j}=1$ for $j=1, \ldots, M$, unless otherwise stated. The minimizing coefficients $\hat{a}_{j}$ then define the reconstructed flip rate through $\hat{F}(x)=\sum_{j=1}^{N} \hat{a}_{j} \phi_{j}(x)$.

An obvious limitation of the projection method is that the method does not converge for non-polynomial $F^{*}$ and polynomial $F^{*}$ with degree $>M$. However, as we shall see in section 3 , the method may still be used to find reasonable approximations in these cases.

```
Algorithm 4 Reconstruction of flip rate coefficients using the \(t\)-method
    1: Require: An integer \(M\), an interval length \(L\), target flip rate function \(F^{*}(x)\), a particle
    speed \(v>0\), a starting position \(-L / 2<x_{0}<L / 2\) and the first \(M\) legendre polynomials
    on ( \(-L / 2, L / 2\) ), \(\phi_{1}, \ldots, \phi_{M}\) (see text for details).
    2: Generate noisy cdfs of the exit time \(W_{1, \text { data }}\left(x_{0}, t\right)\) and \(W_{2, \text { data }}\left(x_{0}, t\right)\) for \(F^{*}(x)\) using Al-
    gorithm 3.
    3: For a given \(\mathbf{a}=\left(a_{0}, a_{1}, \ldots, a_{M-1}\right) \in \mathbb{R}^{M}\), define \(F_{M}(x)=\sum_{j=0}^{M-1} a_{j} \phi_{j}(x)\). Let
    \(W_{1,2}\left(x_{0}, t ; \mathbf{a}\right)\) be the solution to the forward problem (5)-(8) with \(F=F_{M}\), calculated
    using Algorithm 2.
    4: \(\operatorname{Find} \mathbf{a}=\hat{\mathbf{a}}\) that minimizes
\[
\begin{equation*}
\Pi_{1}(\mathbf{a})=\int_{0}^{L / v}\left|W_{1}\left(x_{0}, t ; \mathbf{a}\right)-W_{1, \text { data }}\left(x_{0}, t\right)\right|^{2} \mathrm{~d} t+\int_{0}^{L / v}\left|W_{2}\left(x_{0}, t ; \mathbf{a}\right)-W_{2, \text { data }}\left(x_{0}, t\right)\right|^{2} \mathrm{~d} t \tag{33}
\end{equation*}
\]

Integrating through discontinuities can be avoided by noting that \(W_{1,2}(x, t)=0\) when \(t<\min \left(\frac{L}{2 v}-\frac{x}{v}, \frac{x}{v}+\frac{L}{2 v}\right)\). The lower limits of integration in (33) can be replaced with \(\left(\frac{L}{2 v}-\frac{x_{0}}{v}\right)^{+}\)when \(0 \leq x_{0}<L / 2\) and \(\left(\frac{x_{0}}{v}+\frac{L}{2 v}\right)^{+}\)when \(-L / 2<x_{0} \leq 0\).
5: Output \(\hat{F}(x) \equiv \sum_{j=0}^{M-1} \hat{a}_{j} \phi_{j}(x)\) as the estimate of the flip rate function for the exit time distributions \(W_{1, \text { data }}\left(x_{0}, t\right)\) and \(W_{2, \text { data }}\left(x_{0}, t\right)\).

\section*{3 Results and Discussion}

\section*{Flip Rate Reconstruction}

We used the projection algorithm discussed in section 2.3 to reconstruct flip rate functions from data generated using Monte Carlo simulation (see section 2.2). In the following discussion, let \(N\) be the number of exit times for each initial velocity \(+v,-v\), so that the total number of exit times is always \(2 N\). We also take the starting position \(x_{0}=0\), interval length \(L=1\) and particle speed \(v=1\) unless otherwise stated.

For Fig. 3, we reconstruct some "structurally simple" smooth functions that have few extrema within ( \(-L / 2, L / 2\) ) and find that the accuracy of the reconstructions improves as the noise in the data decreases. For the " \(N=\infty\) " cases, artificial, noiseless data is generated by solving the forward problems eqs. (5)-(8) and (9)-(11) with \(F=F^{*}\). Panels (a-f) indicate that for a given \(N\), the \(t\)-method generally outperforms the \(s\)-method since the associated errors are smaller. In (a,b) we reconstruct a cubic polynomial by recovering \(M=4\) Legendre coefficients. In (c,d) we attempt to reconstruct a transcendental function by representing \(F^{*}(x)\) with \(M=5\) coefficients. Although \(\left\|F_{M}-F^{*}\right\|_{\infty} \nrightarrow 0\) as the noise decreases, we are still able to find a reasonable approximation \(F_{M}\) so that \(\left\|F_{M}-F^{*}\right\|_{\infty}\) is not too large. The \(s\) method converges to the correct solution for perfect data but the inclusion of a small amount

Algorithm 5 Reconstruction of flip rate coefficients using the \(s\)-method
1: Require: An integer \(M\), an interval length \(L\), a target flip rate function \(F^{*}(x)\), a particle speed \(v>0\) and a starting position \(-L / 2<x_{0}<L / 2\) and the first \(M\) Legendre polynomial on ( \(-L / 2, L / 2\) ), \(\phi_{1}, \ldots, \phi_{M}\) (see text for details).
2: Use \(F^{*}(x)\) to generate Laplace-transformed exit time pdfs \(\tilde{w}_{1, \text { data }}\left(x_{0}, s\right)\) and \(\tilde{w}_{2 \text {,data }}\left(x_{0}, s\right)\) through Algorithm 3.
3: For a given \(\mathbf{a} \in \mathbb{R}^{M}\), let \(\tilde{w}_{1,2}\left(x_{0}, s ; \mathbf{a}\right)\) be the solution to the forward problem (9)-(11) calculated using Algorithm 1 with flip rate function defined by \(\mathbf{a}=\left\{a_{0}, a_{2}, \ldots, a_{M-1}\right\}\) :
\[
F_{M}(x)=\sum_{j=0}^{M-1} a_{j} \phi_{j}(x)
\]

4: Find \(\mathbf{a}=\hat{\mathbf{a}}\) that minimizes
\[
\begin{equation*}
\Pi_{2}(\mathbf{a})=\int_{0}^{\infty}\left|\tilde{w}_{1}\left(x_{0}, s ; \mathbf{a}\right)-\tilde{w}_{1, \text { data }}\left(x_{0}, s\right)\right|^{2} \mathrm{~d} s+\int_{0}^{\infty}\left|\tilde{w}_{2}\left(x_{0}, s ; \mathbf{a}\right)-\tilde{w}_{2, \text { data }}\left(x_{0}, s\right)\right|^{2} \mathrm{~d} s \tag{34}
\end{equation*}
\]

The integral in (34) is calculated using using a change of variable \(s[\xi]=\xi /(1-\xi)\) so that
\[
\Pi_{2}=\sum_{j=1}^{2} \int_{0}^{1}\left[\tilde{w}_{j}\left(x_{0}, s[\xi], \mathbf{a}\right)-\tilde{w}_{j, \text { data }}\left(x_{0}, s[\xi]\right)\right]^{2} \frac{\mathrm{~d} \xi}{(1-\xi)^{2}},
\]
which can be computed using the trapezium rule on equally spaced abscissae on \([0,1]\).
5: Output \(\hat{F}(x) \equiv \sum_{j=0}^{M-1} \hat{a}_{j} \phi_{j}(x)\) as the estimate of the flip rate function for the exit time distributions \(\tilde{w}_{1}\left(x_{0}, s\right)\) and \(\tilde{w}_{2}\left(x_{0}, s\right)\).


Figure 3: Reconstructed approximations to flip rate functions \(F^{*}(x)\) from noisy exit time data. (a,b) \(F^{*}(x)=1-0.7 x-0.3 x^{2}+6 x^{3}, M=4 ;(\mathrm{c}, \mathrm{d}) F^{*}(x)=x^{2} e^{-x} / 10+1 / 10, M=5\); \((\mathrm{e}, \mathrm{f}) F^{*}(x)=1+x+3 x^{2}, M=3\). The \(t\)-method was used in the left panels and the \(s\) method was used in the right panels. \(N=\infty\) corresponds to perfect, noiseless data, which is generated by solving the forward problems (5)-(8) and (9)-(11).
of noise renders the method unstable, resulting in a large error. This kind of behavior also occurs with the \(t\)-method when \(M \gtrsim 5\) and is typical in many ill-posed problems (see below). In (e) and (f), we reconstruct flip rate functions from a relatively small number of exit times by taking \(M=3\) basis functions; however, smaller \(M\) restricts the range of admissible target functions.

If we have no a-priori knowledge on \(F^{*}(x)\) (e.g. it may be a high-degree polynomial, have many extrema or be discontinuous), our method may not capture \(F^{*}(x)\) accurately. For our method to be successful, it is important that we know beforehand that \(F^{*}(x)\) is smooth and structurally not too complex. Increasing the number of basis functions \(M\) increases the range of functions we can accurately represent. Providing \(F^{*}\) is smooth enough, it can always be represented through its Taylor series and our method strives to capture its first \(M\) coefficients. Ideally, we would like \(M\) to be as large as possible to represent any
\begin{tabular}{|c|c|c|}
\hline\(M\) & \(\kappa_{t}\) & \(\kappa_{s}\) \\
\hline 2 & \(5.3 \times 10^{2}\) & \(3.1 \times 10^{2}\) \\
3 & \(1.0 \times 10^{4}\) & \(3.3 \times 10^{4}\) \\
4 & \(1.4 \times 10^{5}\) & \(3.5 \times 10^{6}\) \\
5 & \(1.0 \times 10^{6}\) & \(1.6 \times 10^{8}\) \\
\hline
\end{tabular}

Table 1: Local condition numbers \(\kappa_{s}\) and \(\kappa_{t}\) for \(F^{*}(x)=1+x\) corresponding to the objective functions (34) and (33) respectively. \(M\) denotes the number of basis functions used in (32).
\begin{tabular}{|c|c|c|}
\hline\(M\) & \(\kappa_{t}\) & \(\kappa_{s}\) \\
\hline 3 & \(5.0 \times 10^{5}\) & \(1.8 \times 10^{5}\) \\
4 & \(3.5 \times 10^{6}\) & \(1.6 \times 10^{7}\) \\
5 & \(6.7 \times 10^{6}\) & \(5.2 \times 10^{8}\) \\
\hline
\end{tabular}

Table 2: Local condition numbers for \(F^{*}(x)=1-x+x^{2}\) : see Table 1 for details.
\(F^{*}(x) \in C^{\infty}(-L / 2, L / 2)\). However in practice it is difficult to reliably reconstruct \(F^{*}(x)\) (even polynomials) when \(M \gtrsim 5\). The reason, which is common with all projection methods [11, 18], is that as the finite-dimensional approximation to \(F^{*}(x)\) improves with \(M \rightarrow \infty\), the method becomes more unstable due to ill-posedness. In this limit, minimizing the objective functions (33), (34) is prone to large errors.

\section*{Instability of Projection Method}

Numerically, the instability discussed above may be quantified by examining the condition number of the objective function near its minimum. Specifically, we study the Hessian (matrix of second partial derivatives) of the objective functions \(\Pi_{1}\) and \(\Pi_{2}\) in (33) and (34) with respect to the coefficients \(a_{j}, j=0, \ldots, M-1\) :
\[
\begin{equation*}
\left.H_{i j}^{(1)} \equiv \frac{\partial^{2} \Pi_{1}}{\partial a_{i} \partial a_{j}}\right|_{a_{i}=a_{i}^{*}, a_{j}=a_{j}^{*}},\left.H_{i j}^{(2)} \equiv \frac{\partial^{2} \Pi_{2}}{\partial a_{i} \partial a_{j}}\right|_{a_{i}=a_{i}^{*}, a_{j}=a_{j}^{*}}, \tag{35}
\end{equation*}
\]
for \(i, j=0, \ldots, M-1\). In (35), \(a_{i}^{*}\) are the target coefficients of a polynomial flip rate function: \(F^{*}(x)=\sum_{i=0}^{M-1} a_{i}^{*} \phi_{i}(x)\). The condition number of a matrix \(A\) is defined as the ratio of its largest eigenvalue to its smallest: \(\kappa=\lambda_{\max }(A) / \lambda_{\min }(A)\). Since the eigenvalues represent the principle curvatures of \(\Pi_{1,2}\) at the point \(\mathbf{a}^{*}\), they are always positive; a very large condition number indicates that \(\Pi_{1}\) or \(\Pi_{2}\) is locally very flat at \(\mathbf{a}=\mathbf{a}^{*}\) and finding \(\mathbf{a}^{*}\) numerically is prone to errors. On the other hand, a moderate-sized condition number indicates only a small difference in curvatures near \(\mathbf{a}^{*}\) and so finding the minimum numerically should not be difficult. Tables 1 and 2 show that both condition numbers for the \(t\) - and \(s\)-methods grow exponentially as the number of basis functions \(M\) increases. For \(M=5\) basis functions, the condition numbers for the \(t\)-method are consistently two orders of magnitude smaller than those for the \(s\)-method. This suggests that fitting to the exit time data directly (as opposed to its Laplace transform) leads to more effective algorithms and better estimates for the flip rate function. This is confirmed in our numerical experiments since occasionally the \(t\)-method is able to recover \(M=5\) coefficients of a quartic polynomial \(F(x)\), but the \(s\)-method is seldom able to do so.


Figure 4: Error in reconstructed flip rate as a function of particle velocity. Each point is an average over 10 sets of \(2 N=20,000\) exit times. Target flip rate \(F^{*}(x)=\frac{9}{4} x^{2}-\frac{3}{4} x+\frac{9}{16}\), corresponding to target coefficients \(\mathbf{a}^{*}=[0.75,-0.375,0.375,0]\). Initial guesses are (a) \(\mathbf{a}=\) \([1,1,1,1]\), (b) \(\mathbf{a}=[0.76,-0.385,0.365,0.01]\). The same exit times were used for both (a) and (b). The difference in the errors is shown in (c).

\section*{Sensitivity of Reconstruction to Advection Speed}

We also explore the accuracy our reconstruction for different advection speeds \(v\), given a fixed number of exit times when \(F^{*}(x)\) is a polynomial of degree \(\leq M\). In Fig. 4, we see that, for both methods, when the velocity is either much less or much larger than unity, the associated error is large. (Although the upper limit of the objective function (33) depends on the value of \(v\) used, we checked that the non-monotonic behavior in \(\left\|F_{M}-F^{*}\right\|_{\infty}\) was not sensitive to the upper limit of integration.)

In practice, there are always two sources of error in the reconstruction of \(F^{*}\) : the first is from noise in the data and the second stems from the minimization procedure itself:
\[
\begin{equation*}
\text { Total error }=\text { error from noise in data }+ \text { error from minimization } . \tag{36}
\end{equation*}
\]

If the minimization of the objective functions (33) or (34) was achieved with zero error, noisy exit times would still produce an error in the reconstructed \(F\). On the other hand, for noiseless data, the flat minima and large condition numbers discussed above would produce an erroneous \(F\) from the minimization. It is hard to separate the two types of error in (36), but some insight can be gained by comparing Figs. 4 (a) and (b) which differ only in the starting values for the coefficients \(a_{0}, \ldots, a_{M-1}\); in particular the exit times for each value of \(v\) for each figure are identical. When we move the initial guess for the coefficients closer to their target values in (b), we greatly reduce the error in minimization since the accuracy of minimization algorithms depends on the quality of the initial guess. Therefore, the error in (b) comes mainly from noise in the data. Since the exit times were identical for (a) and (b), the difference of the errors in (a) and (b) - shown in (c) - represents the error from minimization which is associated with large condition numbers and flat extrema. We note that the error from minimization from the \(s\)-method is much larger than the corresponding error from the \(t\)-method for a wide range of \(v\) values.

When the dominant error stems from noise in the data (as is the case in Fig. 4(b)), we can understand why \(v=O(1)\) provides the most accurate reconstruction by analyzing how well the Monte Carlo simulations approximate the moments of the exit time distribution. We prove

Theorem 2. Let \(T_{1}^{(n)}(x)\) and \(T_{2}^{(n)}(x)\) be the \(n\)th moments of the exit time conditioned on the particle starting at position \(x\) with initial velocity \(+v\) and \(-v\) respectively. Then the moments have the asymptotic behavior
\[
T_{1,2}^{(k)}(x)= \begin{cases}O\left(\frac{k!}{v^{2 k}}\right), & v \ll 1  \tag{37}\\ O\left(\frac{k!}{v^{k}}\right), & v \gg 1\end{cases}
\]

Proof. We have \(T_{1,2}^{(n)}(x)=\left.(-\partial / \partial s)^{n} \tilde{w}_{1,2}(x, s)\right|_{s=0}\) for \(n \geq 0\), and from eqs. (9) and (10), these moments satisfy the coupled equations
\[
\begin{aligned}
-v \frac{\mathrm{~d} T_{1}^{(n)}}{\mathrm{d} x}-F(x)\left(T_{2}^{(n)}-T_{1}^{(n)}\right) & =n T_{1}^{(n-1)} \\
v \frac{\mathrm{~d} T_{2}^{(n)}}{\mathrm{d} x}-F(x)\left(T_{1}^{(n)}-T_{2}^{(n)}\right) & =n T_{2}^{(n-1)}
\end{aligned}
\]
subject to the boundary conditions \(T_{1}^{(n)}(L / 2)=0\) and \(T_{2}^{(n)}(-L / 2)=0\) where \(n \geq 1\) and \(T_{1,2}^{(0)}(x)=1\). After some algebra, we find expressions for the moments in terms of indefinite integrals:
\[
\begin{align*}
T_{1}^{(n)}(x) & =-\frac{n}{v^{2}} \int \mathrm{~d} x F(x) \int \mathrm{d} x\left[T_{1}^{(n-1)}(x)+T_{2}^{(n-1)}(x)\right]-\frac{n}{v} \int \mathrm{~d} x T_{1}^{(n-1)}(x),  \tag{38}\\
T_{2}^{(n)}(x) & =-\frac{n}{v^{2}} \int \mathrm{~d} x F(x) \int \mathrm{d} x\left[T_{1}^{(n-1)}(x)+T_{2}^{(n-1)}(x)\right]+\frac{n}{v} \int \mathrm{~d} x T_{2}^{(n-1)}(x) . \tag{39}
\end{align*}
\]

When \(v \ll 1\), we retain the first integral in each of eqs. (38) and (39) to find to obtain \(T_{1,2}^{(k)}=O\left(k!/ v^{2 k}\right)\). If \(v \gg 1\), we retain the second integral to find \(T_{1,2}^{(k)}=O\left(k!/ v^{k}\right)\).

In (37), we see that the moments have a different asymptotic form depending on whether \(v\) is small or large. When \(v \ll 1\), the random walker is the diffusive limit where all the moments (except for the zeroth moment) diverge. On the other hand, when \(v \gg 1\), the particle is in the ballistic limit: all moments except for the zeroth moment are asymptotically small and, to leading order, independent of \(F(x)\).

Now consider approximating \(w_{1}\left(x_{0}, t\right)\) or \(w_{2}\left(x_{0}, t\right)\) with their noisy counterparts generated by the Monte-Carlo simulations. How well are the \(w_{1,2}\left(x_{0}, t\right)\) approximated? One way to quantify the accuracy is by calculating the error in the moments of the noisy distribution. Given an initial velocity \(+v\), let \(\left\{\tau_{j}\right\}, 1 \leq j \leq N\) be the \(N\) generated exit times (the following
argument with initial velocity \(-v\) is almost identical). Then by the central limit theorem, the \(k\) th moment is approximately distributed according to
\[
\frac{1}{N} \sum_{j=1}^{N} \tau_{j}^{k} \sim \mathcal{N}\left(T_{1}^{(k)}, \frac{\operatorname{Var}\left(\tau_{j}^{k}\right)}{N}\right)
\]
where \(\mathcal{N}(\mu, \nu)\) is a normal distribution with mean \(\mu\) and variance \(\nu\). Therefore a measure of the error incurred when calculating the \(k\) th sample moment is
\[
N^{-1 / 2} \sqrt{\operatorname{Var}\left(\tau_{j}^{k}\right)}=N^{-1 / 2} \sqrt{T_{1}^{(2 k)}-T_{1}^{(k)^{2}}}=\left\{\begin{align*}
O\left(N^{-1 / 2} v^{-2 k}\right), & v \ll 1,  \tag{40}\\
O\left(N^{-1 / 2} v^{-k}\right), & v \gg 1,
\end{align*}\right.
\]
using (37) and \(\operatorname{Var}\left(\tau_{j}^{k}\right)=E\left[\left(\tau_{j}^{k}-T_{1}^{(k)}\right)^{2}\right]\). Therefore from (40), a quick rule-of-thumb for the accuracy of the Monte-Carlo generated exit time distribution is that the error scales as \(N^{-1 / 2}\) where \(N\) is the number of trials.

It is evident from (40) that for a fixed number of realizations, the error in the \(k\) th moment diverges as \(v^{-2 k}\) as \(v \rightarrow 0\) and the underlying exit time distribution is badly approximated in the limit of small \(v\). On the other hand, as \(v \rightarrow \infty\), although the error in the moments tend to zero, the moments themselves also tend to zero. From eq. (28) the probability that a Broadwell particle with initial velocity \(+v\) exits in time \(t_{c}\) tends to 1 as \(v \rightarrow \infty\) : for large \(v\), the generated list of exit times is populated almost exclusively by \(t_{c}\) (and \(t_{c} \rightarrow 0\) as \(v \rightarrow \infty\) ). From a single exit time it is very difficult to infer any information about \(F(x)\). In both limiting cases, since the distribution of exit times is poorly captured by a finite number of realizations, the quality of the reconstruction suffers.

Finally, we systematically explore the effect of noise on the reconstruction quality. In Fig. 5, we plot the error of the reconstructed \(F_{M}(x)\) against the number of exits. For a wide range of polynomials \(F^{*}(x)\), using both the \(t\) - and \(s\)-methods, we find that the error in the reconstructed function scales as \(O\left(N^{-1 / 2}\right)\). In particular, we see that for \(N=O\left(10^{4}\right)\), the error \(\left\|F_{M}-F^{*}\right\|_{\infty}=O\left(10^{-1}\right)\) whereas \(N\) must exceed \(O\left(10^{6}\right)\) for the error to fall below \(O\left(10^{-2}\right)\). These estimates are mean values: the accuracy resulting from fitting one data set to the next will always vary because the noise in each set is different.

\section*{4 Conclusions}

In this paper, we made three contributions. The first is a pair of algorithms, Algorithms 4 and 5 , that can be used to estimate the flip rate function of a 1 D , constant-speed Broadwell process from the distribution of exit times out of a finite interval. In particular, the \(t\)-method is based on a novel series solution of the backward equation (5)-(8); see Theorem 1. The second is a simulation method, Algorithm 3, that is used in this paper to generate the exit times of a Broadwell particle. The algorithm can accommodate spatially dependent flip rates and velocities. Our final contribution is a set of calculations and asymptotic results that quantify the errors in approximating the exit time distribution with simulated data, and the corresponding error in the flip rate reconstruction.


Figure 5: Dependence of error in reconstructed flip rate function \(F(x)\) on the number of exit times per initial particle state \(N\) when \(F(x)=1-0.7 x-0.3 x^{2}+6 x^{3}\) using (a) \(t\)-method (b) \(s\)-method. For each \(N\), the error is calculated by performing minimizations for 10 data sets and taking the average, with each set containing \(2 N\) exit times.

We found that polynomial transition rates could be reconstructed if the degree of the polynomial was not too large \((\lesssim 4)\) and that fitting to the exit time distribution \((t\)-method) directly versus fitting to the Laplace-transformed distribution ( \(s\)-method) generally allowed the reconstruction of one extra coefficient in the representation of \(F(x)\). Providing our initial guess for the coefficients of \(F(x)\) was not too far from the target coefficients, we were able to find \(F(x)\) to within \(O\left(10^{-1}\right)\) using \(O\left(10^{4}\right)\) exit times. We were also able to find good approximations to non-polynomial flip-rate functions providing they are smooth and slowly varying. Finally, we experimented with reconstructions using different advection speeds. We found that \(v=O(1)\) yielded the most accurate reconstructions because very small or large values of \(v\) in the Monte-Carlo simulations gave poor representations of the true underlying exit time distribution.

Our results suggest that the \(t\)-method is an effective method to infer the spatiallydependent flip rate function of a two-state Broadwell process, if it is known a-priori that this function is smooth and structurally simple. The \(t\)-method involves explicitly solving for the cumulative density functions (5)-(8), tracking the discontinuities via (12)-(13) and minimizing the objective function (33). With this method, one can often find \(M=4\) coefficients from about \(2 N=20,000\) exit times. The \(s\)-method usually reconstructs one less coefficient than the \(t\)-method for the same number of exit times, and is more sensitive to the initial guess. However, it is much simpler to implement and only involves solving the ordinary differential equations (9)-(11) and minimizing (34).

We see two main extensions to this work. The first is to reconstruct spatially dependent advection velocities \(v(x)\) as well as transition rates \(F(x)\). The second is to develop alternative algorithms for reconstruction. We showed in this paper that as the number of coefficients
representing the flip rate function increases, our method becomes unstable due to the presence of flat minima in the objective functions (see eqs. (33) and (34) and Tables 1 and 2). This instability could be alleviated by introducing a small regularization parameter in the objective functions (33), (34) or developing iterative algorithms based directly on (1)-(3) and (5)-(8).

Acknowledgements PWF thanks Rakesh and Tobin Driscoll for helpful discussions. TC was supported by the Army Research Office through grant 58386MA, and the National Science Foundation through grant DMS-1021818. PWF and QH were supported by a University of Delaware Research Foundation (UDRF) grant.

\section*{A Derivation of a multi-state Broadwell process}

A one-dimensional generalized Broadwell model describes a particle that can take any one of \(K\) states. Initially, the particle is at position \(x\) and in state \(i\). A particle in state \(1 \leq k \leq K\) advects within an interval ( \(-L / 2, L / 2\) ) with a velocity function \(v_{k}(y)\) that is single signed on \(-L / 2 \leq y \leq L / 2\) with \(y\) being the current position. While advecting, the particle may transition from state \(i\) to any other state \(j\) with probability \(F_{j i}(y) \mathrm{d} t\) within time interval \((t, t+\mathrm{d} t)\). Also \(F_{j i}(y)\) are positive functions when \(i \neq j\), but \(F_{i i}(y) \equiv 0\). The goal of this appendix is to find the backward equation for the exit time distribution for such a process.


Figure 6: The transitions relevant for computing the Backward equation for the probability density of a multi-state persistent random walk. The quantities \(x, x+\xi\) and \(y\) represent initial, intermediate and final positions while \(i, j\) and \(k\) represent initial, intermediate and final states.

If \(P(y, k, t \mid x, i, 0) \mathrm{d} y\) is the probability that the particle lies between \(y\) and \(y+\mathrm{d} y\) and is in state \(k\) at time \(t\), given that it had started at position \(x\) in state \(i\) at time \(t=0\), the Chapman-Kolmogorov equation [13] is
\[
\begin{equation*}
P(y, k, t+\mathrm{d} t \mid x, i, 0)=\sum_{j=1}^{K} \int_{V_{1}}^{V_{2}} \mathrm{~d} t \quad P(y, k, t+\mathrm{d} t \mid x+\xi, j, \mathrm{~d} t) P(x+\xi, j, \mathrm{~d} t \mid x, i, 0) \mathrm{d} \xi, \tag{41}
\end{equation*}
\]
where \(V_{1}=\min _{k} \min _{-L / 2<y<L / 2} v_{k}(y), V_{2}=\max _{k} \max _{-L / 2<y<L / 2} v_{k}(y)\) and state \(k\) and position \(y\) of the particle at time \(t+\mathrm{d} t\) arise from accounting for all transitions from intermediate
states \(j, j=1, \ldots, K\) and positions \(x+\xi\) at time \(\mathrm{d} t\) : see Fig. 6. Note that \(V_{1}\) can be negative and the integration over \(V_{1} \mathrm{~d} t<\xi<V_{2} \mathrm{~d} t\) corresponds to all possible displacements from position \(x\).

We will assume that the advection velocities and flip rates are explicitly time-independent so that the process is time-homogeneous. Therefore, the probabilities are time-translationally invariant and \(P(y, k, t+\mathrm{d} t \mid x+\xi, j, \mathrm{~d} t)=P(y, k, t \mid x+\xi, j, 0)\). In the following calculations, we make frequent use of delta distributions so Taylor expansions and associated derivatives are to be interpreted in the weak sense.

The probability of transition from \((x, i)\) to \((x+\xi, j)\) in time \(\mathrm{d} t\) can be decomposed into two terms corresponding to continued particle advection when no state flips occur in time \(\mathrm{d} t\), or a state change occuring within \((0, \mathrm{~d} t)\) :
\(P(x+\xi, j, \mathrm{~d} t \mid x, i, 0)=\delta\left(\xi-v_{i}(x) \mathrm{d} t\right)\left(1-\sum_{\substack{\ell=1 \\ \neq i}}^{K} F_{\ell i}(x) \mathrm{d} t\right) \delta_{i j}+\delta(\xi) F_{j i}(x)\left(1-\delta_{i j}\right) \mathrm{d} t+O\left(\mathrm{~d} t^{2}\right)\).
The symbols \(\delta_{i j}\) and \(\delta(\cdot)\) are the usual Kronecker tensor and Dirac delta function, respectively. The first term in Eq. 42 represents the probability that the particle did not transition out of state \(i\) in \((0, \mathrm{~d} t)\), while the second term describes the transition probability from state \(i\) to state \(j \neq i\).

In (42), when \(i=j\), the term proportional to \(F_{j i}\) is zero (no state transitions have occurred) and the probability density originally centered at \(x\) is simply advected to \(x+v_{i}(x) \mathrm{d} t\). Therefore, the probability density of being in position \(x+\xi\) and state \(j=i\) at time \(\mathrm{d} t\) is \(\delta\left(\xi-v_{i}(x) \mathrm{d} t\right)\left(1-\sum F_{\ell i}(x) \mathrm{d} t\right)\). When \(i \neq j\), the term proportional to \(1-\sum F_{\ell i} \mathrm{~d} t\) is zero. To \(O(\mathrm{~d} t)\), we can ignore advection and simply assume the particle changed state in its current position. \({ }^{1}\) Therefore the probability of being at position \(x+\xi\) and state \(j\) at time \(\mathrm{d} t\) is \(\delta(\xi) F_{j i}(x) \mathrm{d} t\).

Upon Taylor-expanding the remaining probability densities in (41):
\[
\begin{equation*}
P(y, k, t+\mathrm{d} t \mid x, i, 0)=P(y, k, t \mid x, i, 0)+\frac{\partial P}{\partial t} \mathrm{~d} t+O\left(\mathrm{~d} t^{2}\right) \tag{43}
\end{equation*}
\]
\[
\begin{equation*}
P(y, k, t \mid x+\xi, j, 0)=P(y, k, t \mid x, j, 0)+\frac{\partial P}{\partial x} \xi+O\left(\xi^{2}\right) \tag{44}
\end{equation*}
\]
we find to order \(\mathrm{d} t\) (note that \(\xi=O(\mathrm{~d} t)\) ),
\[
\begin{equation*}
\frac{\partial}{\partial t} P_{k i}(y, t \mid x, 0)=\sum_{j=1}^{K} L_{j i} P_{k j}(y, t \mid x, 0) \tag{45}
\end{equation*}
\]

\footnotetext{
\({ }^{1}\) If advection was included, \(\delta(\xi) F_{j i} \mathrm{~d} t\) in (42) would be replaced with \(\delta(\xi+O(\mathrm{~d} t)) F_{j i} \mathrm{~d} t\) corresponding to an additional displacement of \(O(\mathrm{~d} t)\) and an associated error of \(O\left(\mathrm{~d} t^{2}\right)\).
}
where
\[
\begin{equation*}
P_{k i}(y, t \mid x, 0) \equiv P(y, k, t \mid x, i, 0), \quad L_{j i}=\delta_{i j} v_{i}(x) \partial_{x}+\left(1-\delta_{i j}\right) F_{j i}-\delta_{i j} \sum_{\substack{\ell=1 \\ \neq i}}^{K} F_{\ell i} . \tag{46}
\end{equation*}
\]

When \(t \leq 0\), the particle is at \(y=x\) in state \(i\) so that
\[
\begin{equation*}
P(y, k, t \mid x, i, 0)=\delta_{i k} \delta(x-y), \quad t \leq 0 \tag{47}
\end{equation*}
\]

Henceforth, we consider the domain \(y \in(-L / 2, L / 2)\) with absorbing boundaries at \(y= \pm L / 2\) so that when \(t>0\) the appropriate boundary conditions are
\[
\begin{align*}
& P(y, k, t \mid x=+L / 2, i, 0)=0, \forall i: v_{i}>0,  \tag{48}\\
& P(y, k, t \mid x=-L / 2, i, 0)=0, \forall i: v_{i}<0 .
\end{align*}
\]

\section*{A. 1 Survival Probabilities}

If we do not distinguish from which boundary the particle eventually exits, we can define the survival probability by integrating Eq. (45) over all positions \(y \in(-L / 2,+L / 2)\) and summing over all possible final states. The survival probability
\[
\begin{equation*}
S_{i}(x, t)=\sum_{k=1}^{K} \int_{-L / 2}^{L / 2} P(y, k, t \mid x, i, 0) \mathrm{d} y \tag{49}
\end{equation*}
\]
describes the probability that a particle started at position \(x \in(-L / 2,+L / 2)\) in state \(i\) has not left through either boundary up to time \(t\). When \(t \leq 0\), we have from (47)
\[
\begin{equation*}
S_{i}(x, t)=1, \quad-L / 2 \leq x \leq L / 2 \tag{50}
\end{equation*}
\]

When \(t>0\), the survival probability obeys
\[
\begin{equation*}
\frac{\partial}{\partial t} S_{i}(x, t)=v_{i}(x) \frac{\partial}{\partial x} S_{i}(x, t)+\sum_{\substack{j=1 \\ \neq i}}^{K} F_{j i}(x) S_{j}(x, t)-\sum_{\substack{\ell=1 \\ \neq i}}^{K} F_{\ell i}(x) S_{i}(x, t) \tag{51}
\end{equation*}
\]
with boundary conditions
\[
\begin{align*}
& S_{i}(L / 2, t)=0, \text { for } i: v_{i}(x)>0, \\
& S_{i}(-L / 2, t)=0 \text { for } i: v_{i}(x)<0 . \tag{52}
\end{align*}
\]

\section*{A. 2 Exit time distributions}

The full exit time distribution is found from the usual definition \(w(t \mid x, i, 0) \equiv w_{i}(x, t)=\) \(-\partial S_{i}(x, t) / \partial t\), so differentiating Eq. (51), we have
\[
\begin{equation*}
\frac{\partial}{\partial t} w_{i}(x, t)=v_{i}(x) \frac{\partial}{\partial x} w_{i}(x, t)+\sum_{\substack{j=1 \\ \neq i}}^{K} F_{j i}(x) w_{j}(x, t)-\sum_{\substack{\ell=1 \\ \neq i}}^{K} F_{\ell i}(x) w_{i}(x, t) \tag{53}
\end{equation*}
\]

The initial condition is found by differentiating (50) in time so that \(w_{i}(x, t)=0\) when \(t \leq 0\). In particular,
\[
\begin{equation*}
w_{i}(x, 0)=0 . \tag{54}
\end{equation*}
\]

From (51) and (52), we have \(S_{i}(L / 2, t)=S_{i}(-L / 2, t)=H(-t)\) for all \(t\). Therefore the boundary conditions for \(w_{i}(x, t)\) are
\[
\begin{align*}
& w_{i}(L / 2, t)=\delta(t), \quad \forall i: v_{i}>0 \\
& w_{i}(-L / 2, t)=\delta(t), \quad \forall i: v_{i}<0 . \tag{55}
\end{align*}
\]

When \(K=2, F_{12}(x)=F_{21}(x)=F(x), v_{1}(x)=-v_{2}(x)=v\) (a positive constant), equations (53)-(55) reduce to (1)-(4).

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